Statistical Methods for Economists

Lecture (7 & 8)c

Three-Way Analysis of Variance (ANOVA) — Latin Squares



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- Three-way ANOVA: Introduction
- Latin squares
- Three-way ANOVA simplified by using Latin squares

We have:

- a set of distinct cars
- a set of distinct drivers
- several types of car-fuel (e.g. fuel with various additives)

We wish to test whether the mileage (the fuel consumption per 100 km) of the car depends also upon the driver who drives the car and on the type of the fuel.



In particular, we have:

- I distinct cars (i = 1, 2, ..., I)
- J distinct drivers (j = 1, 2, ..., J)
- K distinct types of fuel (k = 1, 2, ..., K)

There are three factors in this example:

- factor A = the car (i = 1, 2, ..., I)
- factor B = the driver (j = 1, 2, ..., J)
- factor C = the type of the fuel (k = 1, 2, ..., K)





Considering the $IJK = I \times J \times K$ distinct combinations of the factors

(the Cartesian product), we assume that each combination is tested n_{ijk} -times. We thus have a sample

$$\begin{array}{l} \mathbf{y}_{ijkl} \\ \text{of the underlying random variables} \\ Y_{ijkl}: \Omega \rightarrow \mathbb{R} \end{array} \qquad \text{for} \quad \begin{cases} \ i = 1, 2, \dots, l \\ j = 1, 2, \dots, J \\ k = 1, 2, \dots, K \\ l = 1, 2, \dots, n_{ijk} \end{cases}$$

These random variables are assumed to be normal $(Y_{ijkl} \sim \mathcal{N}(\mu_{ijk}, \sigma^2))$,

independent (uncorrelated) and homoscedastic (with the same variance σ^2).



We assume that the effect of the factors A, B, C is additive.

Moreover, it is possible to distinguish many combinations of interactions, e.g.:

1. No interactions between / among the factors:

 $Y_{ijkl} \approx \mu + \alpha_i + \beta_j + \gamma_k$

2. Interactions between the factors – <u>interactions of the 1st order</u>: $Y_{ijkl} \approx \mu + \alpha_i + \beta_j + \gamma_k + \lambda_{ij}^{AB} + \lambda_{ik}^{AC} + \lambda_{jk}^{BC}$

3. Interactions among the factors – interactions of the 2nd order:

$$Y_{ijkl} \approx \mu + \alpha_i + \beta_j + \gamma_k + \lambda_{ijk}^{ABC}$$

4. All interactions between and among the factors:

i = 1, 2, ..., Ij = 1, 2, ..., Jk = 1, 2, ..., K $l = 1, 2, ..., n_{ijk}$



We assume that the effect of the factors A, B, C is additive.

Moreover, it is possible to distinguish many combinations of interactions, e.g.:

1.
$$Y_{ijkl} = \mu + \alpha_i + \beta_j + \gamma_k + \varepsilon_{ijkl}$$

2.
$$Y_{ijkl} = \mu + \alpha_i + \beta_j + \gamma_k + \lambda_{ij}^{AB} + \lambda_{ik}^{AC} + \lambda_{jk}^{BC} + \varepsilon_{ijkl}$$

3.
$$Y_{ijkl} = \mu + \alpha_i + \beta_j + \gamma_k + \lambda_{ijk}^{ABC} + \varepsilon_{ijkl}$$

4.
$$Y_{ijkl} = \mu + \alpha_l + \beta_j + \gamma_k + \lambda_{ij}^{AB} + \lambda_{ik}^{AC} + \lambda_{jk}^{BC} + \lambda_{ijk}^{ABC} + \varepsilon_{ijkl}$$

 $\begin{aligned} \varepsilon_{ijkl} &\sim \mathcal{N}(0, \sigma^2) \\ i &= 1, 2, ..., l \\ j &= 1, 2, ..., J \\ k &= 1, 2, ..., K \\ l &= 1, 2, ..., n_{ijk} \end{aligned}$

In the models, such as

$$Y_{ijkl} = \mu + \alpha_i + \beta_j + \gamma_k + \lambda_{ij}^{AB} + \lambda_{ik}^{AC} + \lambda_{jk}^{BC} + \lambda_{ijk}^{ABC} + \varepsilon_{ijkl}$$

the parameters

$$\mu, \alpha_i, \beta_j, \gamma_k, \lambda_{ij}^{\text{AB}}, \lambda_{ik}^{\text{AC}}, \lambda_{jk}^{\text{BC}}, \lambda_{ijk}^{\text{ABC}} \in \mathbb{R}$$

are <u>unknown</u>.

We assume that the unknown parameters are normalized so that:

$$\sum_{i=1}^{I} \alpha_i = 0 \qquad \qquad \sum_{j=1}^{J} \beta_j = 0 \qquad \qquad \sum_{k=1}^{K} \gamma_k = 0$$



No.

We assume that the unknown parameters are normalized so that:

$$\sum_{i=1}^{I} \lambda_{ij}^{AB} = 0 = \sum_{j=1}^{J} \lambda_{ij}^{AB} \qquad \sum_{i=1}^{I} \lambda_{ik}^{AC} = 0 = \sum_{k=1}^{K} \lambda_{ik}^{AC} \qquad \sum_{j=1}^{J} \lambda_{jk}^{BC} = 0 = \sum_{k=1}^{K} \lambda_{jk}^{BC}$$

and
$$\sum_{k=1}^{K} \lambda_{ijk}^{ABC} = 0 \qquad \sum_{j=1}^{J} \lambda_{ijk}^{ABC} = 0 \qquad \sum_{i=1}^{I} \lambda_{ijk}^{ABC} = 0$$

or
$$\sum_{i=1}^{I} \sum_{j=1}^{J} \lambda_{ijk}^{ABC} = 0 \qquad \sum_{i=1}^{I} \sum_{k=1}^{K} \lambda_{ijk}^{ABC} = 0 \qquad \sum_{j=1}^{J} \sum_{k=1}^{K} \lambda_{ijk}^{ABC} = 0$$



It is possible to consider the general situation with a general number $n_{ijk} \ge 1$ of observations for each combination of the factors and it is also possible to formulate and test various null hypotheses $(\alpha_i = 0 / \lambda_{ij}^{AB} = 0 / \lambda_{ijk}^{ABC} = 0 / \text{etc.})$, but there are plenty of calculations and the resulting formulas are complicated.

This is why we shall study the following special case only:

(see the next slide)



We assume for simplicity that the number of the levels of all three factors is the same:

$$I = J = K$$

In addition to that, we do not perform all the observations for each combination of the factors (because the number of the necessary experiments could soon become infeasible).

Instead, we perform either exactly one observation $(n_{ijk} = 1)$ or no observation

at all $(n_{ijk} = 0)$ according to the scheme called Latin square.

Latin squares in Three-way ANOVA





A Latin square of order *N* is an arrangement of *N* symbols, such as $\{1, 2, ..., N\}$, where each symbol is repeated *N*-times, into an $N \times N$ square table in such a way that

- · in each column, each symbol occurs exactly once and
- in each row, each symbol occurs exactly once.

For example:

1 2 3 2 3 1 3 1 2 1 2 3 3 1 2 2 3 1





The name "*Latin square*" is inspired by the work of **Leonhard Euler** (1707–1783), a Swiss mathematician, physicist, astronomer, geographer, logician, and engineer who used the upper-case letters of the Latin alphabet as the symbols in the square:

														A	В	С	D	E
						Α	В	С	D	Α	В	С	D	В	С	D	Е	A
 A	В	С	Α	В	С	В	С	D	А	С	D	А	В	C	D	Е	А	В
В	С	А	С	А	В	С	D	Α	В	D	С	В	А	D	Е	А	В	C
С	А	В	В	С	А	D	А	В	С	В	А	D	С	E	А	В	С	D



Assume that each of the three factors A, B, C has the same number of levels

$$I = J = K = N$$

for some natural number $N \ge 2$.

- Arrange (or denote) the factors A, B, C so that
 - we assume that factors A and B do have some effect

on the observed values

 we ask (test the hypothesis) whether factor C has any effect on the observed values



- Arrange the N symbols $\{1, 2, ..., N\}$ into a Latin square of order N.
- There are plenty of distinct Latin squares of order N.
- It is recommended: The Latin square should be chosen randomly.

Now, given the Latin square of type $N \times N$, define the numbers n_{lfk} as follows:

$$n_{ijk} = \begin{cases} 1, & \text{if the symbol } k \text{ is at the position } (i,j), \\ 0, & \text{otherwise.} \end{cases} \text{ for } \begin{cases} i = 1, 2, \dots, N \\ j = 1, 2, \dots, N \\ k = 1, 2, \dots, N \end{cases}$$



Then,

- for each i = 1, 2, ..., N and for each j = 1, 2, ..., N,
- find the unique $k \in \{1, 2, ..., N\}$ such that $n_{ijk} = 1$,
- set up Factor A to the level *i* and set up Factor B to the level *j*,
- set up Factor C to the level k, and carry out the experiment,
- observe the numerical outcome y_{ijk1} of the random variable Y_{ijk1}

Recall that Factors A and B are assumed to have some effect. We ask (test the hypothesis) whether Factor C has any effect.

We assume

- that the effect of Factors A, B, C is additive and
- that there are no interactions between / among the factors:

$$Y_{ijk1} = \mu + \alpha_i + \beta_j + \gamma_k + \varepsilon_{ijk1}$$

for
$$\begin{cases} i = 1, 2, ..., N \\ j = 1, 2, ..., N \\ k = 1, 2, ..., N \\ \& n_{ijk} = 1 \end{cases}$$

where the meaning of the (unknown) parameters $\mu, \alpha_i, \beta_j, \gamma_k \in \mathbb{R}$ is as follows:

- μ the common mean value
- α_i the effect of the level *i* of Factor A (for i = 1, 2, ..., N)
- β_j the effect of the level *j* of Factor B (for j = 1, 2, ..., N)
- γ_k the effect of the level k of Factor C (for k = 1, 2, ..., N)



We assume

- that the effect of Factors A, B, C is additive and
- that there are no interactions between / among the factors:

$$Y_{ijk1} = \mu + \alpha_i + \beta_j + \gamma_k + \varepsilon_{ijk1}$$

for
$$\begin{cases} i = 1, 2, ..., N \\ j = 1, 2, ..., N \\ k = 1, 2, ..., N \\ \& n_{ijk} = 1 \end{cases}$$

Moreover, we assume that the (unknown) parameters $\alpha_i, \beta_j, \gamma_k \in \mathbb{R}$ are normalized so that:

$$\sum_{i=1}^{N} \alpha_i = 0 \qquad \qquad \sum_{j=1}^{N} \beta_j = 0 \qquad \qquad \sum_{k=1}^{N} \gamma_k = 0$$



We assume

- that the effect of Factors A, B, C is additive and
- that there are no interactions between / among the factors:

$$Y_{ijk1} = \mu + \alpha_i + \beta_j + \gamma_k + \varepsilon_{ijk1}$$

We assume

- Factor A has some effect (that is $\alpha_i \neq 0$ for some $i \in \{1, 2, ..., N\}$)
- Factor B has some effect (that is $\beta_j \neq 0$ for some $j \in \{1, 2 ..., N\}$)

We ask - test the null hypothesis - whether Factor C has any effect:

$$H_{\rm C}: \quad \gamma_1 = \gamma_2 = \cdots = \gamma_N = 0$$



for
$$\begin{cases} i = 1, 2, ..., N \\ j = 1, 2, ..., N \\ k = 1, 2, ..., N \\ \& n_{ijk} = 1 \end{cases}$$

Three-Way ANOVA without interactions and simplified by using Latin squares

 $Y_{ijk1} = \mu + \alpha_i + \beta_j + \gamma_k + \varepsilon_{ijk}$



Assume that we have a sample

of observations of the random variables

$$Y_{ijk1} = \mu + \alpha_i + \beta_j + \gamma_k + \varepsilon_{ijk1}$$

Yijk1

for
$$\begin{cases} i = 1, 2, ..., N \\ j = 1, 2, ..., N \\ k = 1, 2, ..., N \\ \& n_{ijk} = 1 \end{cases}$$

where $\mu, \alpha_i, \beta_j, \gamma_k \in \mathbb{R}$ are fixed (but <u>unknown</u>) parameters normalized

so that

$$\sum_{i=1}^{N} \alpha_i = 0 \qquad \qquad \sum_{j=1}^{N} \beta_j = 0 \qquad \qquad \sum_{k=1}^{N} \gamma_k = 0$$



Assume that we have a sample

of observations of the random variables

$$Y_{ijk1} = \mu + \alpha_i + \beta_j + \gamma_k + \varepsilon_{ijk1}$$

Yijk1

for
$$\begin{cases} i = 1, 2, ..., N \\ j = 1, 2, ..., N \\ k = 1, 2, ..., N \\ \& n_{ijk} = 1 \end{cases}$$

...and

$$\varepsilon_{ijk1} \sim \mathcal{N}(0,\sigma^2)$$

are mutually independent random variables with the same variance $\sigma^2 \in \mathbb{R}^+$ (the variance σ^2 is also unknown).



Denote the index set:

$$\mathcal{S} = \left\{ (i, j, k, 1) : i, j, k \in \{1, 2, \dots, N\}, n_{ijk} = 1 \right\}$$

Notice that there are exactly N^2 elements in the index set S.

Indeed, for each pair (i, j) of the indices for i = 1, 2, ..., N and for j = 1, 2, ..., N, there exists exactly one $k \in \{1, 2, ..., N\}$ such that $n_{ijk} = 1$.



Stack the observations y_{ijk1} into the N²-dimensional vector

$$y = (y_{ijk1})_{(i,j,k,1) \in \mathcal{S}} \in \mathbb{R}^{\mathcal{S}}$$

and introduce the sample mean:

$$\bar{y} = \frac{1}{N^2} \sum_{i=1}^{N} \sum_{j=1}^{N} \sum_{\substack{k=1\\n_{ijk}=1}}^{N} y_{ijk1}$$

This sample mean is an estimate of the parameter μ (the common mean value):

 $\bar{y} \approx \mu$



Let $\mathbf{1} = (1)_{(i,j,k,1) \in S} \in \mathbb{R}^{S}$ be the vector of N^{2} ones and introduce the line

 H_0 :

$$L = \{ \mathbf{1}\lambda : \lambda \in \mathbb{R} \} =$$
$$= \{ \mathbf{z} \in \mathbb{R}^{\mathcal{S}} : z_{ijk1} = \mu, \ \mu \in \mathbb{R} \}$$

which corresponds to the null hypothesis that

$$Y_{ijk1} = \mu + \varepsilon_{ijk1} \qquad \text{for} \begin{cases} j = 1, 2, ..., N \\ k = 1, 2, ..., N \\ \& n_{ijk} = 1 \end{cases}$$

that is

$$\alpha_1 = \cdots = \alpha_N = 0$$
 $\beta_1 = \cdots = \beta_N = 0$ $\gamma_1 = \cdots = \gamma_N = 0$

(cf. one-way ANOVA)

 $i = 1, 2, \dots, N$

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Moreover, introduce the subspace

$$H_{\mathrm{A}} = \left\{ \boldsymbol{z} \in \mathbb{R}^{\mathcal{S}} : \boldsymbol{z}_{ijk1} = \boldsymbol{\mu} + \beta_j + \gamma_k, \ \boldsymbol{\mu}, \beta_j, \gamma_k \in \mathbb{R}, \ \sum_{j=1}^N \beta_j = \sum_{k=1}^N \gamma_k = 0 \right\}$$

which corresponds to the null hypothesisi = 1, 2, ..., N H_A : $Y_{ijk1} = \mu + \beta_j + \gamma_k + \varepsilon_{ijk1}$ fori = 1, 2, ..., Nk = 1, 2, ..., N

$$\alpha_1 = \cdots = \alpha_N = 0$$

Observe that the line

 $L \subset H_A$

Introduce also the subspace

$$H_{\mathrm{B}} = \left\{ z \in \mathbb{R}^{\mathcal{S}} : z_{ijk1} = \mu + \alpha_i + \gamma_k, \ \mu, \alpha_i, \gamma_k \in \mathbb{R}, \ \sum_{j=1}^N \alpha_i = \sum_{k=1}^N \gamma_k = 0 \right\}$$

which corresponds to the null hypothesis $H_{\mathbf{B}}: \quad Y_{ijk1} = \mu + \alpha_i + \gamma_k + \varepsilon_{ijk1}$ for $\begin{cases} i = 1, 2, ..., N \\ j = 1, 2, ..., N \\ k = 1, 2, ..., N \\ k = 1, 2, ..., N \\ k = n_{ijk} = 1 \end{cases}$ that is $\beta_1 = \cdots = \beta_N = 0$

Observe that the line

 $L \subset H_{\rm B}$



And introduce also the subspace

$$H_{\mathsf{C}} = \left\{ z \in \mathbb{R}^{\mathcal{S}} : z_{ijk1} = \mu + \alpha_i + \beta_j, \ \mu, \alpha_i, \beta_j \in \mathbb{R}, \ \sum_{j=1}^N \alpha_i = \sum_{k=1}^N \beta_j = 0 \right\}$$

which corresponds to the null hypothesis

$$H_{C}$$
: $Y_{ijk1} = \mu + \alpha_{i} + \beta_{j} + \varepsilon_{ijk1}$ for $\begin{cases} i = 1, 2, ..., N \\ j = 1, 2, ..., N \\ k = 1, 2, ..., N \\ k = 1, 2, ..., N \\ \& n_{ijk} = 1 \end{cases}$
that is

$$\gamma_1=\cdots=\gamma_N=0$$

Observe that the line

 $L \subset H_{C}$



Finally, introduce the subspace

$$M = \begin{cases} \mathbf{z} \in \mathbb{R}^{\mathcal{S}} : z_{ijk1} = \mu + \alpha_i + \beta_j + \gamma_k, \quad \mu, \alpha_i, \beta_j, \gamma_k \in \mathbb{R}, \\ \sum_{i=1}^N \alpha_i = \sum_{j=1}^N \beta_j = \sum_{k=1}^N \gamma_k = 0 \end{cases}$$

which corresponds to the model under consideration:

$$Y_{ijk1} = \mu + \alpha_i + \beta_j + \gamma_k + \varepsilon_{ijk1} \quad \text{for} \quad \begin{cases} i = 1, 2, ..., N \\ j = 1, 2, ..., N \\ k = 1, 2, ..., N \\ k = 1, 2, ..., N \\ \& n_{ijk} = 1 \end{cases}$$

Three-way ANOVA with no interactions: Dimensions

Notice that the dimension of

- the line

is

$$L = \left\{ z \in \mathbb{R}^{S} : z_{ijk1} = \mu, \ \mu \in \mathbb{R} \right\}$$
1

— the subspace

$$H_{A} = \left\{ z \in \mathbb{R}^{S} : z_{ijk1} = \mu + \beta_{j} + \gamma_{k}, \ \mu, \beta_{j}, \gamma_{k} \in \mathbb{R}, \ \sum_{j=1}^{N} \beta_{j} = \sum_{k=1}^{N} \gamma_{k} = 0 \right\}$$
is

$$(1 + N + N) - 1 - 1 = 2N - 1$$





Notice that the dimension of

- the subspace

$$H_{B} = \left\{ z \in \mathbb{R}^{S} : z_{ijk1} = \mu + \alpha_{i} + \gamma_{k}, \ \mu, \alpha_{i}, \gamma_{k} \in \mathbb{R}, \ \sum_{j=1}^{N} \alpha_{i} = \sum_{k=1}^{N} \gamma_{k} = 0 \right\}$$
is
$$(1 + N + N) - 1 - 1 = 2N - 1$$

— the subspace

$$H_{C} = \left\{ z \in \mathbb{R}^{S} : z_{ijk1} = \mu + \alpha_{i} + \beta_{j}, \ \mu, \alpha_{i}, \beta_{j} \in \mathbb{R}, \ \sum_{j=1}^{N} \alpha_{i} = \sum_{k=1}^{N} \beta_{j} = 0 \right\}$$
is

$$(1+N+N)-1-1 = 2N-1$$



Notice that the dimension of

- the subspace $M = \begin{cases} \mathbf{z} \in \mathbb{R}^{S} : z_{ijk1} = \mu + \alpha_i + \beta_j + \gamma_k, \quad \mu, \alpha_i, \beta_j, \gamma_k \in \mathbb{R}, \\ \sum_{i=1}^{N} \alpha_i = \sum_{j=1}^{N} \beta_j = \sum_{k=1}^{N} \gamma_k = 0 \end{cases}$ is (1 + N + N + N) - 1 - 1 - 1 = 3N - 2



Letting $\hat{y}_{ijk1} = \mu - \alpha_i - \beta_j - \gamma_k$, solve the Least Squares Problem:





Letting $\hat{y}_{Aijk1} = \mu - \beta_j - \gamma_k$, $\hat{y}_{Bijk1} = \mu - \alpha_i - \gamma_k$, $\hat{y}_{Cijk1} = \mu - \alpha_i - \beta_j$, solve also the problems:

$$\min_{\widehat{y}_{A}\in H_{A}}\sum_{i=1}^{N}\sum_{j=1}^{N}\sum_{\substack{k=1\\n_{ijk}=1}}^{N}(y_{ijk1}-\mu-\beta_{j}-\gamma_{k})^{2} = \min_{\widehat{y}_{A}\in H_{A}}||y-\widehat{y}_{A}||^{2}$$

and

$$\min_{\hat{\mathbf{y}}_{B}\in H_{B}} \sum_{i=1}^{N} \sum_{j=1}^{N} \sum_{\substack{k=1\\n_{ijk}=1}}^{N} (y_{ijk1} - \mu - \alpha_{i} - \gamma_{k})^{2} = \min_{\hat{\mathbf{y}}_{B}\in H_{B}} \|\mathbf{y} - \hat{\mathbf{y}}_{B}\|^{2}$$



and

$$\min_{\widehat{\mathbf{y}}_{C}\in H_{C}}\sum_{i=1}^{N}\sum_{j=1}^{N}\sum_{\substack{k=1\\n_{ijk}=1}}^{N}(y_{ijk1}-\mu-\alpha_{i}-\beta_{j})^{2} = \min_{\widehat{\mathbf{y}}_{C}\in H_{C}}||\mathbf{y}-\widehat{\mathbf{y}}_{C}||^{2}$$

Finally, letting $\bar{y} = \mu$, solve the Least Squares Problem:

$$\min_{1\bar{y}\in L}\sum_{i=1}^{N}\sum_{j=1}^{N}\sum_{\substack{k=1\\n_{ijk}=1}}^{N}(y_{ijk1}-\mu)^{2} = \min_{1\bar{y}\in L}||y-1\bar{y}||^{2}$$
We have and denote:

$$\underbrace{\|\boldsymbol{y} - \mathbf{1}\bar{\boldsymbol{y}}\|^2}_{SS_{TOTAL}} = \underbrace{\|\widehat{\boldsymbol{y}} - \widehat{\boldsymbol{y}}_A\|^2}_{SS_A} + \underbrace{\|\widehat{\boldsymbol{y}} - \widehat{\boldsymbol{y}}_B\|^2}_{SS_B} + \underbrace{\|\widehat{\boldsymbol{y}} - \widehat{\boldsymbol{y}}_C\|^2}_{SS_C} + \underbrace{\|\boldsymbol{e}\|^2}_{RSS}$$

where, recall, we have:

 $\hat{y}_{\mathrm{A}ijk1} = \hat{\mu} + \hat{\beta}_j + \hat{\gamma}_k$

$$\begin{aligned} \hat{y}_{ijk1} &= \hat{\mu} + \hat{\alpha}_i + \hat{\beta}_j + \hat{\gamma}_k \\ \hat{y}_{Bijk1} &= \hat{\mu} + \hat{\alpha}_i + \hat{\gamma}_k \\ \bar{y} &= \hat{\mu} + \hat{\alpha}_j + \hat{\beta}_j \end{aligned}$$



$$SS_{TOTAL} = \|y - 1\bar{y}\|^2 = \sum_{i=1}^{N} \sum_{j=1}^{N} \sum_{\substack{k=1 \ n_{ijk}=1}}^{N} (y_{ijk1} - \bar{y})^2 = \sum_{i=1}^{N} \sum_{\substack{j=1 \ n_{ijk}=1}}^{N} \sum_{\substack{k=1 \ n_{ijk}=1}}^{N} (y_{ijk1} - \hat{\mu})^2$$

$$RSS = \|\boldsymbol{e}\|^2 = \|\boldsymbol{y} - \hat{\boldsymbol{y}}\|^2 = \sum_{i=1}^N \sum_{j=1}^N \sum_{\substack{k=1 \ n_{ijk}=1}}^N (y_{ijk1} - \hat{y}_{ijk1})^2 =$$

$$= \sum_{i=1}^{N} \sum_{j=1}^{N} \sum_{\substack{k=1 \ n_{ijk}=1}}^{N} (y_{ijk1} - \hat{\mu} - \hat{\alpha}_i - \hat{\beta}_j - \hat{\gamma}_k)^2$$



Three-way ANOVA with no interactions

$$SS_{A} = \|\widehat{\mathbf{y}} - \widehat{\mathbf{y}}_{A}\|^{2} = \sum_{i=1}^{N} \sum_{j=1}^{N} \sum_{\substack{k=1 \ n_{ijk}=1}}^{N} (\widehat{y}_{ijk1} - \widehat{y}_{Aijk1})^{2} =$$
$$= \sum_{i=1}^{N} \sum_{j=1}^{N} \sum_{\substack{k=1 \ n_{ijk}=1}}^{N} (\widehat{\mu} + \widehat{\alpha}_{i} + \widehat{\beta}_{j} + \widehat{\gamma}_{k} - \widehat{\mu} - \widehat{\beta}_{j} - \widehat{\gamma}_{k})^{2} =$$
$$= \sum_{i=1}^{N} \sum_{j=1}^{N} \sum_{\substack{k=1 \ n_{ijk}=1}}^{N} \widehat{\alpha}_{i}^{2} = N \sum_{i=1}^{N} \widehat{\alpha}_{i}^{2}$$



Three-way ANOVA with no interactions

$$SS_{B} = \|\widehat{\mathbf{y}} - \widehat{\mathbf{y}}_{B}\|^{2} = \sum_{i=1}^{N} \sum_{j=1}^{N} \sum_{\substack{k=1 \ n_{ijk}=1}}^{N} (\widehat{y}_{ijk1} - \widehat{y}_{Bijk1})^{2} =$$
$$= \sum_{i=1}^{N} \sum_{j=1}^{N} \sum_{\substack{k=1 \ n_{ijk}=1}}^{N} (\widehat{\mu} + \widehat{\alpha}_{i} + \widehat{\beta}_{j} + \widehat{\gamma}_{k} - \widehat{\mu} - \widehat{\alpha}_{i} - \widehat{\gamma}_{k})^{2} =$$
$$= \sum_{i=1}^{N} \sum_{j=1}^{N} \sum_{\substack{k=1 \ n_{ijk}=1}}^{N} (\widehat{\mu} + \widehat{\alpha}_{i} + \widehat{\beta}_{j} + \widehat{\gamma}_{k} - \widehat{\mu} - \widehat{\alpha}_{i} - \widehat{\gamma}_{k})^{2} =$$



$$SS_{C} = \|\widehat{\mathbf{y}} - \widehat{\mathbf{y}}_{C}\|^{2} = \sum_{i=1}^{N} \sum_{j=1}^{N} \sum_{\substack{k=1 \ n_{ijk}=1}}^{N} (\widehat{y}_{ijk1} - \widehat{y}_{Cijk1})^{2} =$$
$$= \sum_{i=1}^{N} \sum_{j=1}^{N} \sum_{\substack{k=1 \ n_{ijk}=1}}^{N} (\widehat{\mu} + \widehat{\alpha}_{i} + \widehat{\beta}_{j} + \widehat{\gamma}_{k} - \widehat{\mu} - \widehat{\alpha}_{i} - \widehat{\beta}_{j})^{2} =$$
$$= \sum_{i=1}^{N} \sum_{j=1}^{N} \sum_{\substack{k=1 \ n_{ijk}=1}}^{N} \widehat{\gamma}_{k}^{2} = N \sum_{k=1}^{N} \widehat{\gamma}_{k}^{2}$$





By solving the above Least Squares Problems

 $(\partial F/\partial \mu = \partial F/\partial \alpha_i = \partial F/\partial \beta_j = \partial F/\partial \gamma_k = 0$ etc.; <u>do it as an exercise</u>), we obtain:

$$\hat{\mu} = \frac{1}{N^2} \sum_{i=1}^{N} \sum_{j=1}^{N} \sum_{\substack{k=1 \ n_{ijk}=1}}^{N} y_{ijk1} = \bar{y} = \bar{y}...$$

$$\hat{\alpha}_{i} = \frac{1}{N} \sum_{j=1}^{N} \sum_{\substack{k=1 \ n_{ijk}=1}}^{N} y_{ijk1} - \hat{\mu} = \bar{y}_{i..} - \bar{y} \quad \text{for } i = 1, 2, ..., N$$



By solving the above Least Squares Problems

 $(\partial F/\partial \mu = \partial F/\partial \alpha_i = \partial F/\partial \beta_j = \partial F/\partial \gamma_k = 0$ etc.; <u>do it as an exercise</u>), we obtain:

$$\hat{\beta}_{j} = \frac{1}{N} \sum_{i=1}^{N} \sum_{\substack{k=1 \\ n_{ijk}=1}}^{N} y_{ijk1} - \hat{\mu} = \bar{y}_{.j.} - \bar{y} \quad \text{for } j = 1, 2, ..., N$$

$$\hat{y}_k = \frac{1}{N} \sum_{i=1}^N \sum_{\substack{j=1\\n_{ijk}=1}}^N y_{ijk1} - \hat{\mu} = \bar{y}_{..k} - \bar{y}$$
 for $k = 1, 2, ..., N$

Put together, we have:

SS_{TOTAL} =
$$\sum_{i=1}^{N} \sum_{j=1}^{N} \sum_{\substack{k=1 \ n_{ijk}=1}}^{N} (y_{ijk1} - \hat{\mu})^2 = \sum_{i=1}^{N} \sum_{\substack{j=1 \ n_{ijk}=1}}^{N} \sum_{\substack{k=1 \ n_{ijk}=1}}^{N} (y_{ijk1} - \bar{y})^2$$

RSS =
$$\sum_{i=1}^{N} \sum_{j=1}^{N} \sum_{\substack{k=1 \ n_{ijk}=1}}^{N} (y_{ijk1} - \hat{\mu} - \hat{\alpha}_i - \hat{\beta}_j - \hat{\gamma}_k)^2 =$$

$$=\sum_{i=1}^{N}\sum_{j=1}^{N}\sum_{\substack{k=1\\n_{ijk}=1}}^{N}\left(y_{ijk1}-\bar{y}_{i..}-\bar{y}_{..k}+2\bar{y}\right)^{2}$$





Remark: The quantity

$$s^2 = \frac{\text{RSS}}{\dim M^{\perp}} = \frac{\text{RSS}}{(N-1)(N-2)}$$

is an estimate of the unknown σ^2 , that is $s^2 \approx \sigma^2$. It holds

$$\mathbf{E}[s^2] = \sigma^2$$

Recall that dim $M^{\perp} = N^2 - \dim M = (N - 1)(N - 2)$ — see the figure above!

Put together, we have:

$$SS_{A} = N \sum_{i=1}^{N} \hat{\alpha}_{i}^{2} = N \sum_{i=1}^{N} (\bar{y}_{i..} - \bar{y})^{2}$$

$$SS_{B} = N \sum_{j=1}^{N} \hat{\beta}_{j}^{2} = N \sum_{j=1}^{N} (\bar{y}_{.j.} - \bar{y})^{2}$$

$$SS_{C} = N \sum_{k=1}^{N} \hat{\gamma}_{k}^{2} = N \sum_{k=1}^{N} (\bar{y}_{..k} - \bar{y})^{2}$$



Three-way ANOVA with no interactions

Recall that it holds:

$$SS_{TOTAL} = SS_A + SS_B + SS_C + RSS$$

$$\sum_{i=1}^{N} \sum_{j=1}^{N} \sum_{\substack{k=1 \ n_{ijk}=1}}^{N} \left(y_{ijk1} - \bar{y} \right)^2 = N \sum_{i=1}^{N} (\bar{y}_{i\cdots} - \bar{y})^2 + N \sum_{j=1}^{N} (\bar{y}_{\cdot j} - \bar{y})^2 + N \sum_{k=1}^{N} (\bar{y}_{\cdot k} - \bar{$$

$$+\sum_{i=1}^{N}\sum_{j=1}^{N}\sum_{\substack{k=1\\n_{ijk}=1}}^{N}(y_{ijk1}-\bar{y}_{i..}-\bar{y}_{.j.}-\bar{y}_{..k}+2\bar{y})^{2}$$





We use the theory of Linear Regression (Theorem 8) to test Hypothesis H_C :

If the null hypothesis

$$H_{\rm C}: \quad \gamma_1 = \gamma_2 = \cdots = \gamma_N = 0$$

holds true, then

$$\frac{SS_{C}}{RSS} / \frac{\dim M - \dim H_{C}}{N^{2} - \dim M} = \frac{SS_{C}}{RSS} / \frac{N-1}{(N-1)(N-2)} \sim F_{N-1,(N-1)(N-2)}$$

Three-way ANOVA with no interactions: Test for $H_{\rm C}$ It holds: M^{\perp} $\cot^2 \varphi = \frac{(\widehat{\boldsymbol{y}} - \widehat{\boldsymbol{y}}_{\rm C})^{\rm T} (\widehat{\boldsymbol{y}} - \widehat{\boldsymbol{y}}_{\rm C})}{\rm RSS}$ (the orthogonal complement = = the space of $e = y - \hat{y}$ the residuals) $(\cot a \, \varphi)^2 / \frac{N-1}{N^2 - 3N + 2} \sim F_{N-1, (N-1)(N-2)}$ subspace of dimension $N^2 - 3N + 2$ Μ subspace of dimension 3N - 2*H*_C subspace of dimension the dimension of its complement within the subspace of dimension 3N - 2 is 2N - 1N-1

• Given the sample y_{ijk1} of the random variables $Y_{ijk1} = \mu + \alpha_i + \beta_j + \gamma_k + \varepsilon_{ijk1}$ where $\mu, \alpha_i, \beta_j, \gamma_k \in \mathbb{R}$ are (unknown) parameters such that $\sum_{i=1}^{N} \alpha_i = \sum_{j=1}^{N} \beta_j = \sum_{k=1}^{N} \gamma_k = 0$ and $\varepsilon_{ijk1} \sim \mathcal{N}(0, \sigma^2)$ are mutually independent random variables for i = 1, 2, ..., N, j = 1, 2, ..., N, k = 1, 2, ..., N such that $n_{ijk} = 1$, formulate the null hypothesis: H_{C} : $\gamma_1 = \gamma_2 = \cdots = \gamma_N = 0$

• The alternative hypothesis is $H_{C1} \equiv \neg H_C$, i.e. $\gamma_k \neq 0$ for some $k \in \{1, 2, ..., N\}$





Calculate the statistic

$$F = \frac{SS_{C}}{RSS} / \frac{DF_{C}}{DF_{RSS}} = \frac{N \sum_{k=1}^{N} (\bar{y}_{..k} - \bar{y})^{2}}{\sum_{i=1}^{N} \sum_{j=1}^{N} \sum_{k=1}^{N} (y_{ijk1} - \bar{y}_{i..} - \bar{y}_{.j.} - \bar{y}_{..k} + 2\bar{y})^{2}} / \frac{N - 1}{N^{2} - 3N + 2}$$

• If the null hypothesis is true, then we have by the Theorem

$$F \sim F_{N-1,(N-1)(N-2)}$$

• Choose the level of significance, a small number $\alpha > 0$, such as $\alpha = 5$ %, other popular values are $\alpha = 10$ % or $\alpha = 1$ % or $\alpha = 0.1$ % etc.



find the critical value

$$c = F_{N-1,(N-1)(N-2)} (1-\alpha)$$

so that $\int_{c}^{+\infty} f(x) dx = \alpha$ where f is the density of the *F*-distribution with N - 1 and (N - 1)(N - 2) degrees of freedom

- if $F \in [c, +\infty)$, the critical region, then <u>reject</u> the null hypothesis
- if $F \in [0, c)$, then <u>do not reject</u> (or <u>fail to reject</u>) the null hypothesis