Statistical Methods for Economists

Lecture 9

Full Factorial Design of Experiments



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- Motivation & Introduction
- Full Factorial Experiments
- Graphical assessment of factor / interaction significance
- Graphs of interaction effects

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- We consider an observed variable Y, such as the quality of a product.
- We also consider several factors A, B, C, D, ..., such as the conditions during the production.
- And we conjecture that the levels of the factors (and the interactions between the factors) influence the response variable *Y*, which we wish to maximize, say.
- The purpose is to identify the significant factors that influence the response variable *Y* most, and to find their optimal levels.
- We study the effect of the factors A, B, C, D, ... on the response variable Y



- We study the effect of the factors A, B, C, D, ... on the response variable Y by means of experiment.
- We consider only a few (a finite number) of distinct values of each of the factors.
- We usually consider only two distinct values of each of the factors.
- The first value is "low" and is denoted by "-".
- The second value is "high" and is denoted by "+".
- If *n* factors are considered,

then there are 2^n various combinations of the levels of the factors.

- If we consider n = 2 factors A and B, then the arrangement of the factor levels is called a 2×2 or 2² factorial design.
- There are 2×2 = 4 factorial points, which constitute a square:







• There are 2×2×2 = 8 factorial points, which constitute a cube:





- The theory of the Full Factorial Experiments is based upon the theory of <u>Multiple Linear Regression</u> (again).
- The theory and all the calculations are simplified significantly if we assume that the number of the experiments carried out in each case is the same.

 In other words, considering n factors and assuming that the experiment is repeated K-times for each combination of the factors, we have

 $2^n \times K$

observations of the experiment in total.

• In other words, considering n factors and assuming that the experiment is repeated *K*-times for each combination of the factors, we have $2^n \times K$

numerical outcomes of the mutually independent random variables

$$Y_{sk} \sim \mathcal{N}(\mu_s, \sigma^2)$$
 for $s \in S$ and for $k = 1, 2, ..., K$

where

$$\mathcal{S} = \{\pm\}^n = \{\pm, -\}^n$$

is the index set of all the 2^n possible combinations of the levels of the factors.

We thus have a sample

$$y_{s1}, y_{s2}, \dots, y_{sK}$$
 for $s \in S$

of the observations of the mutually independent random variables

 $Y_{s1}, Y_{s2}, \dots, Y_{sK} \sim \mathcal{N}(\mu_s, \sigma^2) \quad \text{for} \quad s \in \mathcal{S}$

where S is the index set of all the 2^n combinations of the levels of the factors.

(Here, the expected values μ_s and the variance σ^2 are unknown. The variance σ^2 is assumed to be the same – <u>homoskedasticity</u>.)



Consider a spring.

The observed variable Y is the lifespan of the spring,

i.e. the number of pressures of the spring until it cracks.

We consider three factors:

- Factor L = the length of the spring
- Factor G = the thickness of the wire of the spring
- Factor T = the material of (the wire of) the spring



We consider the following levels of the three factors:

• Factor L = the length of the spring

"-" or "-1" = 10 cm

"+" or "+1" = 15 cm

Factor G = the thickness of the wire of the spring
 "-" or "-1" = 5 mm

"+" or "+1" = 7 mm

- Factor T = the material of (the wire of) the spring
 "-" or "-1" = Material / alloy "A"
 - "+" or "+1" = Material / alloy "B"



The experiment was carried out 2× for each combination of the factors:



	Factor	Observed value					
L	G	т	$Y_{\pm\pm\pm1}$	$Y_{\pm\pm\pm2}$			
-	—	_	77	81			
+	_	_	98	96			
-	+	_	76	74			
+	+	_	90	94			
-	—	+	63	65			
+	—	+	82	86			
-	÷	÷	72	74			
+	+	+	92	88			



Now, the goal is to identify the factors, including their interactions, with significant effect on the lifespan of the spring, i.e. the observed values of the variable Y.

We shall use the theory of Multiple Linear Regression to this end.

We assume the model with the intercept term:

 $Y = 1\beta_0 + x_L\beta_L + x_G\beta_G + x_T\beta_T + x_{LG}\beta_{LG} + x_{LT}\beta_{LT} + x_{GT}\beta_{GT} + x_{LGT}\beta_{LGT}$



We assume the model with the intercept term:

 $Y = 1\beta_0 + x_{\rm L}\beta_{\rm L} + x_{\rm G}\beta_{\rm G} + x_{\rm T}\beta_{\rm T} + x_{\rm LG}\beta_{\rm LG} + x_{\rm LT}\beta_{\rm LT} + x_{\rm GT}\beta_{\rm GT} + x_{\rm LGT}\beta_{\rm LGT}$

The effect of the factor or the interaction is <u>significant</u> if and only if the respective coefficient β is non-zero.

→ We shall use the *t*-test for the respective parameter β (see Theorem 6) to this end.

We set up the matrix X as the first step:



"—" = "—1	" &	"+" = "+1"		Interactio	n = the p	roduct of	the Factor	S.
Intercept		Factor			Intera	action		
0	L	G	Т	LG	LT	GT	LGT	
1	—	-	—	÷	Ŧ	÷	_	
1	+			-		÷	+	
1	—	+		—	+	—	+	
1	+	+	—	÷	_	—	_	
1	-	—	÷	÷	—	—	+	
1	+	-	÷	_	÷	—	_	
1	-	+	4	—	-	÷	_	
1	÷	+	÷	+	÷	+	+	

This is a $2^n \times 2^n$ matrix in general.

Since we carried out each experiment 2× for each combination of the factors, we double each row of the table:

Intercept Observed Factor Interaction 0 G Т LG LT GT LGT $Y_{\pm\pm\pm k}$ **X**: 1 -1 -1 -1 **y**: +1 +1 +1 -1 77 1 -1 -1 -1 +1 +1 +1 -1 81 1 -1 +1 -1 -1 -1 +1 +1 98 1 +1 -1 +1 -1 -1 -1 +1 96 1 -1 +1 -1 -1 +1 -1 +1 76 1 -1 +1 -1 -1 +1 -1 +1 74 1 +1 +1 -1 -1 -1 -1 +1 90 1 +1 +1 -1 +1 -1 -1 -1 94 1 -1 -1 -1 +1 +1 -1 +1 63 1 -1 -1 +1 +1 -1 +1 -1 65 1 +1 -1 +1 +1 -1 -1 -1 82 1 +1 -1 -1 -1 +1 +1 -1 86 1 -1 +1 +1 -1 -1 +1 -1 72 -1 1 +1 -1 -1 +1 -1 +1 74 1 +1 +1 +1 +1 +1 +1 +1 92 +1 +1 +1 +1 +1 +1 +1 88

Interaction = the product of the Factors.

This is a $(2^n \times K) \times (2^n)$ matrix in general.



We shall perform the *t*-test for each individual parameter β .

Recall the Corollary of Theorem 6:

If $\beta_j = 0$, then

$$T = \frac{b_j}{\sqrt{s^2}\sqrt{c_{jj}}} \sim t_{N-\text{rank}(X)}$$

where

- b_f is the estimate of the parameter β_f
- s^2 is the residual variance = mean square error
- c_{jj} the *j*-th element on the diagonal of the matrix $C = (X^T X)^{-1}$
- N is the total number of observations, we have $N = 2^n \times K$ here





Having recalled the Corollary of Theorem 6:

if
$$\beta_j = 0$$
 then $T = \frac{b_j}{\sqrt{s^2}\sqrt{c_{jj}}} \sim t_{N-\text{rank}(X)}$

- the only task is to calculate everything in this special case, i.e. for this matrix X.

Calculating, we obtain that:

$$\boldsymbol{X}^{\mathrm{T}}\boldsymbol{X} = \begin{pmatrix} \boldsymbol{N} & & & \\ & \boldsymbol{N} & & \\ & & \ddots & \\ & & & & \boldsymbol{N} \end{pmatrix}$$

i.e. the $2^n \times 2^n$ diagonal matrix with the number $N = 2^n \times K$ on the diagonal.

In our example, we have n = 3 factors (L, G, T),

and each experiment is replicated K = 2 times. We thus have

 $N=2^n\times K=16$

Moreover, the result $X^T X$ is a $2^n \times 2^n$, i.e. 8×8 , matrix

with the number 16 on its diagonal:

$$\boldsymbol{X}^{\mathrm{T}}\boldsymbol{X} = \begin{pmatrix} 16 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 \\ 0 & 16 & 0 & 0 & 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 16 & 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 16 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 16 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & 16 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & 0 & 16 & 0 \\ 0 & 0 & 0 & 0 & 0 & 0 & 0 & 16 & 0 \\ \end{pmatrix}$$





Calculating, we obtain that:

$$\boldsymbol{C} = (\boldsymbol{X}^{\mathrm{T}}\boldsymbol{X})^{-1} = \begin{pmatrix} 1/N & & \\ & 1/N & \\ & \ddots & \\ & & \ddots & \\ & & & 1/N \end{pmatrix}$$

i.e. the $2^n \times 2^n$ diagonal matrix with the number $c_{jj} = 1/N = 1/(2^n \times K)$ on the diagonal.

Moreover, the estimates are

$$\boldsymbol{b} = \boldsymbol{C}\boldsymbol{X}^{\mathrm{T}}\boldsymbol{y} = \left(\boldsymbol{X}^{\mathrm{T}}\boldsymbol{X}\right)^{-1}\boldsymbol{X}^{\mathrm{T}}\boldsymbol{y} = \frac{1}{N}\boldsymbol{X}^{\mathrm{T}}\boldsymbol{y}$$

N.

In our example, we

$$\boldsymbol{C} = \left(\boldsymbol{X}^{\mathrm{T}}\boldsymbol{X}\right)^{-1} = \begin{pmatrix} 1/16 & 0 & 0 & 0 & 0 & 0 & 0 & 0 \\ 0 & 1/16 & 0 & 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 1/16 & 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 1/16 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 1/16 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & 1/16 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & 0 & 1/16 & 0 \\ 0 & 0 & 0 & 0 & 0 & 0 & 0 & 1/16 \end{pmatrix}$$

And the estimates are:

- the intercept term = the sample mean of the observed values:

$$b_0 = \frac{1}{N} (X_0)^T y = \frac{1}{16} (+77 + 98 + 76 + 90 + 63 + 82 + 72 + 92 + 81 + 96 + 74 + 94 + 65 + 86 + 74 + 88) = 1308/16 = 81.75$$



The estimates of the <u>effects of the factors</u> are:

$$b_{\rm L} = \frac{1}{N} (X_{\rm L})^{\rm T} y = \frac{1}{16} (-77 - 98 + 76 + 90 - 63 - 82 + 72 + 92 - 81 - 96 + 74 + 94 - 65 - 86 + 74 + 88) = 144/16 = 9$$

$$b_{\rm G} = \frac{1}{N} (X_{\rm G})^{\rm T} y = \frac{1}{16} (-77 - 98 - 76 - 90 + 63 + 82 + 72 + 92 - 81 - 96 - 74 - 94 + 65 + 86 + 74 + 88) = 12/16 = 0.75$$

$$b_{\rm T} = \frac{1}{N} (X_{\rm T})^{\rm T} y = \frac{1}{16} (-77 - 98 - 76 - 90 - 63 - 82 - 72 - 92 + 81 + 96 + 74 + 94 + 65 + 86 + 74 + 88) = -64/16 = -4$$

The estimates of the effects of the interactions between / among the factors are:

$$b_{LG} = \frac{1}{N} (X_{LG})^T y = \frac{1}{16} (+77 + 98 - 76 - 90 - 63 - 82 + 72 + 92 + 81 + 96 - 74 - 94 - 65 - 86 + 74 + 88) = -8/16 = -0.5$$

$$b_{\text{LT}} = \frac{1}{N} (X_{\text{LT}})^{\text{T}} y = \frac{1}{16} (+77 + 98 - 76 - 90 + 63 + 82 - 72 - 92 - 81 - 96 + 74 + 94 - 65 - 86 + 74 + 88) = 4/16 = 0.25$$

$$b_{\rm GT} = \frac{1}{N} (X_{\rm GT})^{\rm T} y = \frac{1}{16} (+77 + 98 + 76 + 90 - 63 - 82 - 72 - 92 - 81 - 96 - 74 - 94 + 65 + 86 + 74 + 88) = 48/16 = 3$$

$$b_{\text{LGT}} = \frac{1}{N} (X_{\text{LGT}})^{\text{T}} y = \frac{1}{16} (+77 + 98 + 76 + 90 - 63 - 82 - 72 - 92 - 81 - 96 - 74 - 94 + 65 + 86 + 74 + 88)$$



We further have to calculate the residual variance = mean square error:

$$s^{2} = \frac{\text{RSS}}{N - \text{rank}(X)} = \frac{e^{T}e}{N - \text{rank}(X)} = \frac{(y - \hat{y})^{T}(y - \hat{y})}{N - \text{rank}(X)}$$

where

$$N = 2^n \times K$$

and it is easy to see that

$$\operatorname{rank}(X) = 2^n$$

Moreover, the predicted values are

$$\widehat{\boldsymbol{y}} = \boldsymbol{X}\boldsymbol{b} = \boldsymbol{X}\boldsymbol{C}\boldsymbol{X}^{\mathsf{T}}\boldsymbol{y} = \frac{1}{N}\boldsymbol{X}\boldsymbol{X}^{\mathsf{T}}\boldsymbol{y}$$

No.

Calculating, we obtain that:

$$\boldsymbol{X}\boldsymbol{X}^{\mathrm{T}} = \begin{pmatrix} 2^{n}\boldsymbol{E} & & \\ & 2^{n}\boldsymbol{E} & \\ & & \ddots & \\ & & & 2^{n}\boldsymbol{E} \end{pmatrix}$$

i.e. the $(2^n \times K) \times (2^n \times K)$ matrix with the $K \times K$ matrix $2^n E$ repeated 2^n -times on its diagonal. The symbol

E is the $K \times K$ matrix of all ones

and

 $2^n E$ is the $K \times K$ matrix of the numbers 2^n



Substituting into the above equation

$$\widehat{y} = Xb = XCX^{T}y = \frac{1}{N}XX^{T}y = \frac{1}{2^{n} \times K}XX^{T}y$$

and calculating, we obtain the predicted values:

$$\hat{y}_{sk} = \frac{1}{K}(y_{s1} + y_{s2} + \dots + y_{sK})$$
 for $s \in S$ for $k = 1, 2, \dots, K$

where - recall -

$$\mathcal{S} = \{\pm\}^n = \{\pm, -\}^n$$

is the index set of all the 2^n possible combinations of the levels of the factors and *K* is the number of the replications of the experiment for the given



We have:

$$s^{2} = \frac{\text{RSS}}{N - \text{rank}(X)} = \frac{e^{T}e}{N - \text{rank}(X)} = \frac{(y - \hat{y})^{T}(y - \hat{y})}{N - \text{rank}(X)} = \frac{(y - \hat{y})^{T}(y - \hat{y})}{2^{n} \times K - 2^{n}}$$

Substituting, we have:

$$s^{2} = \frac{(y - \hat{y})^{T}(y - \hat{y})}{(K - 1)2^{n}} = \frac{1}{(K - 1)2^{n}} \sum_{s \in S} \sum_{k=1}^{K} (y_{sk} - \hat{y}_{sk})^{2} =$$
$$= \frac{1}{(K - 1)2^{n}} \sum_{s \in S} \sum_{k=1}^{K} \left(y_{sk} - \frac{y_{s1} + y_{s2} + \dots + y_{sK}}{K} \right)^{2}$$



In our example, we have:

	L	-	-	+	+	-	-	+	+	-	-	+	+	-	-	+	+
Factor	G	-	-	-	-	+	+	+	+	-	-	-	-	+	+	+	+
	Т	-	-	-	-	-	-	-	-	+	+	+	+	+	+	+	+
Observed	$y_{\pm\pm\pm k}$	77	81	98	96	76	74	90	94	63	65	82	86	72	74	92	88
Sample means	$\hat{y}_{\pm\pm\pm}$	7	9	9	7	7	5	9	2	6	4	8	4	7	3	9	0

 $RSS = (77 - 79)^{2} + (81 - 79)^{2} + (98 - 97)^{2} + (96 - 97)^{2} + (76 - 75)^{2} + (74 - 75)^{2} + (90 - 92)^{2} + (94 - 92)^{2} + (63 - 64)^{2} + (65 - 64)^{2} + (82 - 84)^{2} + (86 - 84)^{2} + (72 - 73)^{2} + (74 - 73)^{2} + (92 - 90)^{2} + (88 - 90)^{2} = 40$

and

$$s^{2} = \frac{\text{RSS}}{(K-1)2^{n}} = \frac{40}{(2-1) \times 2^{3}} = \frac{40}{8} = 5$$

Full Factorial Experiments

Finally, we calculate the statistic

$$T = \frac{b_j}{\sqrt{s^2}\sqrt{c_{jj}}} =$$

$$=\frac{\frac{1}{N}(X_j)^{\mathrm{T}} \mathbf{y}}{\sqrt{\frac{RSS}{(K-1)2^n}}\sqrt{\frac{1}{N}}} = \sqrt{\frac{1}{K2^n}(X_j)^{\mathrm{T}} \mathbf{y}}\sqrt{\frac{(K-1)2^n}{RSS}} = \sqrt{\frac{(K-1)/K}{RSS}} (X_j)^{\mathrm{T}} \mathbf{y}$$





Finally, we perform the *t*-test of the significance of the factor or the interaction.

To this end, we compare the value of the above statistic T with the quantile

$$t_{N-2^n}\left(1-\frac{\alpha}{2}\right)$$

• If $|T| \ge t_{N-2^n}(1-\alpha/2)$, the critical region,

then **<u>reject</u>** the null hypothesis that $b_j = 0$,

that is, consider the factor / interaction significant.

• If $|T| < t_{N-2^n}(1 - \alpha/2)$, then <u>fall to reject</u> the null hypothesis $b_j = 0$, that is, consider the factor / interaction insignificant – <u>neglect it</u>.



In our example – choosing significance level $\alpha = 5 \%$ – we have the quantile

$$q = t_{N-2^n} \left(1 - \frac{\alpha}{2} \right) = t_{16-8} \left(1 - \frac{0.05}{2} \right) = t_8(0.975) \doteq 2.306004$$

and

 $b_{\rm L}$:

 $b_{\rm G}$:

 $T = \frac{b_{\rm L}}{\sqrt{s^2}\sqrt{c_{\rm LL}}} = \frac{9}{\sqrt{5}\sqrt{1/16}} \doteq 16.100 \ge q \implies \text{significant}$ $T = \frac{b_{\rm G}}{\sqrt{s^2}\sqrt{c_{\rm GG}}} = \frac{0.75}{\sqrt{5}\sqrt{1/16}} \doteq 1.342 < q \implies \text{neglect}$



$$T = \frac{b_{\rm T}}{\sqrt{s^2}\sqrt{c_{\rm TT}}} = \frac{-4}{\sqrt{5}\sqrt{1/16}} \doteq -7.155 \le -q \quad \Rightarrow \text{ significant}$$

$$b_{\rm LG}:$$

$$T = \frac{b_{\rm LG}}{\sqrt{s^2}\sqrt{c_{\rm LGLG}}} = \frac{-0.5}{\sqrt{5}\sqrt{1/16}} \doteq -0.894 > -q \quad \Rightarrow \text{ neglect}$$

$$b_{\rm LT}:$$

$$T = \frac{b_{\rm LT}}{\sqrt{s^2}\sqrt{c_{\rm LTLT}}} = \frac{0.25}{\sqrt{5}\sqrt{1/16}} \doteq 0.447 < q \quad \Rightarrow \text{ neglect}$$

$$b_{\rm GT}:$$

$$T = \frac{b_{\rm GT}}{\sqrt{s^2}\sqrt{c_{\rm GTGT}}} = \frac{3}{\sqrt{5}\sqrt{1/16}} \doteq 5.367 > q \quad \Rightarrow \text{ significant}$$



$$T = \frac{b_{\rm LGT}}{\sqrt{s^2}\sqrt{c_{\rm LGTLGT}}} = \frac{-0.25}{\sqrt{5}\sqrt{1/16}} \doteq -0.447 > -q \implies \text{neglect}$$

CONCLUSION:

b_{LGT}:

the factors L T and the interaction GT have significant effect on the lifespan of the spring.

Moreover, we conclude that

$$Y \approx b_0 + x_L b_L + x_T b_T + x_{GT} b_{GT}$$

 $\approx 81.75 + 9x_L - 4x_T + 3x_{GT}$ where $x_L, x_T, x_{GT} \in \{-1, +1\}$

Graphical assessment of factor / interaction significance



If the number of the replications of each experiment is

K = 1

then the above computational procedure cannot be used

<u>because</u>

$$s^2 = \frac{RSS}{(K-1)2^n}$$

that is, we cannot perform the calculations because we would divide by zero, or we would always calculate T = 0.

We perform the graphical assessment of the significance then.





Assuming K = 1, so that $N = 2^n$,

the graphical assessment proceeds as follows:

- Sort all the estimates b_s , for $s \in S$, in ascending order.
- Denote the *i*-th least value by $b_{(i)}$ for $i = 1, 2, ..., 2^n 1$ so that we have

$$b_{(1)} \le b_{(2)} \le \dots \le b_{(2^{n}-1)}$$

Calculate the values

$$P_i = \frac{i - \frac{1}{2}}{2^n - 1}$$
 for $i = 1, 2, ..., 2^n - 1$

Plot the points

$$[b_{(i)}, P_i]$$
 for $i = 1, 2, ..., 2^n - 1$

In our example, we have:

$$b_{\rm T} < b_{\rm LG} < b_{
m LGT} < b_{
m LT} < b_{
m G} < b_{
m GT} < b_{
m L}$$

 $-4 < -0.5 < -0.25 < 0.25 < 0.75 < 3 < 9$

and P_i :

$$\frac{0.5}{7} < \frac{1.5}{7} < \frac{2.5}{7} < \frac{3.5}{7} < \frac{4.5}{7} < \frac{5.5}{7} < \frac{6.5}{7}$$







The points $[b_{(i)}, P_i]$ take the shape of an S-curve sometimes.

Those points that lie out of the approximately linear middle part of the S-curve indicate the significant factors or interactions.

In our example: the points

 $b_{\rm T}$ $b_{\rm GT}$ $b_{\rm L}$

lie out of the approximately linear middle part of the S-curve,

which is the same result that we obtained by using the t-test above.

Graphs of interaction effects





There is an interaction effect between two factors in a full factorial experiment

when the effect of one independent variable (factor) depends on the level

of another independent variable (factor).

Considering two factors A and B, take

- all the observations when factors A and B were at the levels --, respectively
- all the observations when factors A and B were at the levels -+, respectively
- all the observations when factors A and B were at the levels +-, respectively
- all the observations when factors A and B were at the levels ++, respectively

and calculate the <u>average</u> (sample mean) for each of the four groups

Full Factorial Experiments: Interaction effect

We then have:



We then depict the values in the form of a chart

(here as the dependence of Factor B on Factor A):

• If the lines are ≈ parallel,

then the interaction is not significant.

• If the lines are not parallel,

then the interaction is significant.







In our example, we have:

	L	-	-	+	+	-	_	+	+	-	-	+	+	-	_	+	+
Factor	G	-	-	-	-	+	+	+	+	-	-	-	-	+	+	+	+
	Т	_	-	-	-	-	-	-	_	+	+	+	+	+	+	+	+
Observed	$y_{\pm\pm\pm k}$	77	81	98	96	76	74	90	94	63	65	82	86	72	74	92	88

	+ -	76,□ 74,□ 72,□ 7	90, 🗆 94, 🗆 92, 🗆 8			63, 0 65, 0 72, 0 7 82, 0 86, 0 92, 0 8			
Footor O		74	91	Contor T	+	68.5	87		
Factor G		77, 🗆 81, 🗆 63, 🗆 6	98, 🗆 96, 🗆 82, 🗆 8	Factor T		77, 🗆 81, 🗆 76, 🗆 7	798,□96,□90,□9		
		71.5	90.5		-	77	94.5		
		-	+			-	+		
		Fact	or L			Fac	tor L		











\approx parallel \Rightarrow the interaction is not significant











\approx parallel \Rightarrow the interaction is not significant











NOT parallel ⇒ the interaction is SIGNIFICANT